

Mixture Model Challenge

Václav Šmíd

April 8, 2022

Recapitulation

- ▶ What is latent variable of mixture model?

Recapitulation

- ▶ What is latent variable of mixture model?
- ▶ What is the conditional likelihood of observation of a mixture model?

Recapitulation

- ▶ What is latent variable of mixture model?
- ▶ What is the conditional likelihood of observation of a mixture model?
- ▶ What is the advantage of one-hot representation?

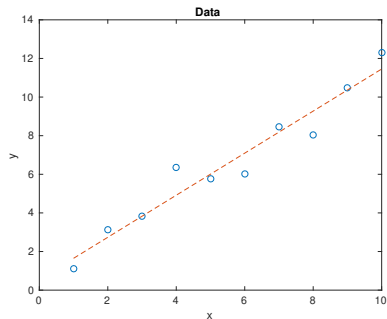
Review of building blocks

Linear regression:

$$y_i = \mathbf{x}_i\theta + e_i,$$

$$e_i \sim \mathcal{N}(0, \sigma_e),$$

$$p(y_i|x_i, \theta) = \mathcal{N}(\mathbf{x}_i\theta, \sigma_e),$$



Review of building blocks

Linear regression likelihood:

$p(\theta_i)$

$$y_i = \mathbf{x}_i\theta + e_i,$$

$$e_i \sim \mathcal{N}(0, \sigma_e),$$

$$p(y_i|x_i, \theta) = \mathcal{N}(\mathbf{x}_i\theta, \sigma_e),$$

Prior (ridge):

$$p(\theta_i) = \mathcal{N}(0, \alpha)$$

Review of building blocks

Linear regression likelihood:

$$y_i = \mathbf{x}_i\theta + e_i,$$

$$e_i \sim \mathcal{N}(0, \sigma_e),$$

$$p(y_i | x_i, \theta) = \mathcal{N}(\mathbf{x}_i\theta, \sigma_e),$$

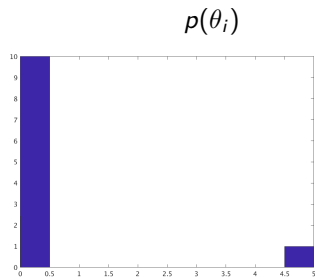
Prior (ridge):

$$p(\theta_i) = \mathcal{N}(0, \alpha)$$

Prior Sparse:

$$\theta = [0, 0, 0, 5, 0, 0, 0, 0, 0, 0, 0]$$

$$p(\theta_i) = \begin{cases} St(\nu, 0, \sigma_0) \\ 0.9\mathcal{N}(0, \sigma_0) + 0.1\mathcal{N}(0, \sigma_1) \\ 0.9\mathcal{N}(0, \sigma_0) + 0.1\mathcal{N}(0, 100\sigma_0) \end{cases}$$



Patlak Rutland plot

Sequence of scintigraphic images of kidneys.

filename: drsprg_023

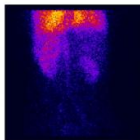
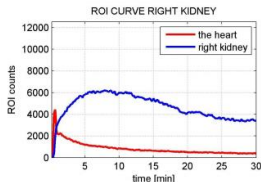
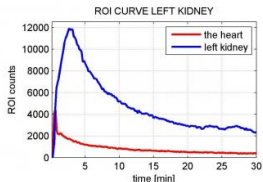
gender = F, age = 31 yrs

CKD stage = 2, LK = 77 %

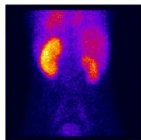
serum Cr - 0, Cr clearance - 0

99mTc-MAG3 - 0, 51Cr-EDTA - 0

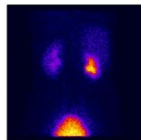
57Co-FLOOD - 1



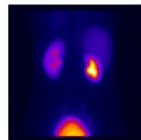
0 - 1 min



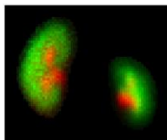
1 - 2 min



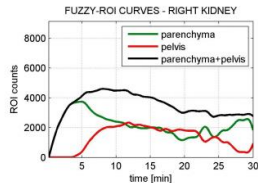
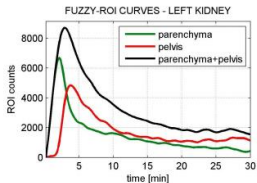
29 - 30 min



MEAN IMAGE



FUZZY ROIS

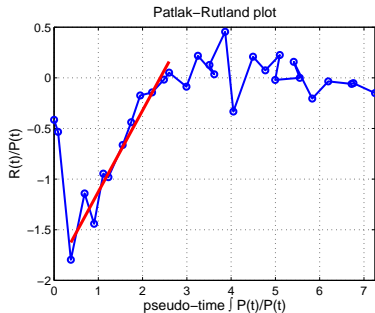


Challenge: Patlak Rutland plot

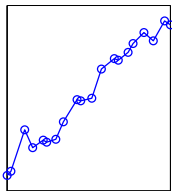
Patlak Rutland plot is a ratio of the parenchyma curve over the cumulative sum of the heart curve.

Diagnostically important is the **slope** of the linear part, which:

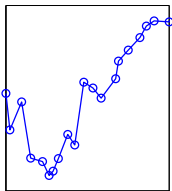
- ▶ typically starts around 1min
- ▶ typically ends around 3min
- ▶ with outliers
- ▶ manual fit is the norm



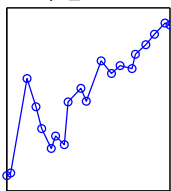
dtpa_1i1c.crv



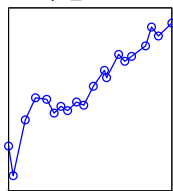
dtpa_2i1c.crv



dtpa_3i1c.crv



dtpa_4i1c.crv



Models?

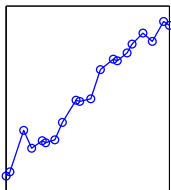
- ▶ no heuristics
 - ▶ probabilistic model providing either:
 - ▶ estimate of \hat{a}
 - ▶ posterior distribution $p(a)$
- where a is the slope of:

$$y_i = ax_i + b + e_i$$

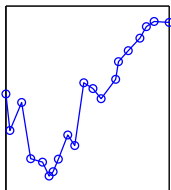
for points belonging to it.



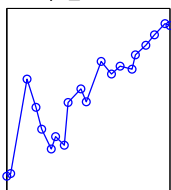
dtpa_1i1c.crv



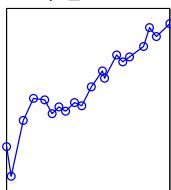
dtpa_2i1c.crv



dtpa_3i1c.crv



dtpa_4i1c.crv

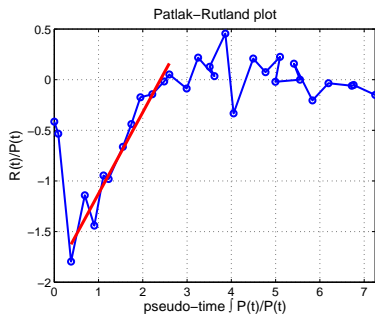


Model 1: Robust regression

Model

$$y_i = ax_i + b + e_i$$

is valid for all $x_i \in [1, 3]$ with heavy-tailed noise $p(e_i)$.



Model 1: Robust regression

Model

$$y_i = ax_i + b + e_i$$

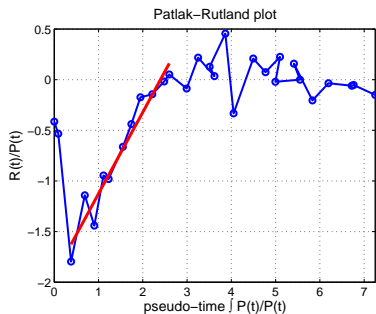
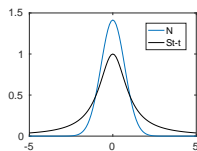
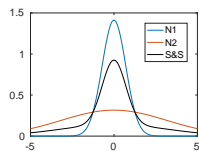
is valid for all $x_i \in [1, 3]$ with heavy-tailed noise $p(e_i)$.

Possibilities:

- ▶ Spike and slab,

$$p(e_i) = \alpha \mathcal{N}(0, v_1) + (1 - \alpha) \mathcal{N}(0, v_2),$$

- ▶ Student-t $p(e_i) = \mathcal{St}(\nu, 0, \sigma)$,



Latent representations

Spike and slab

$$p(e_i | v_1, v_2) = \alpha \mathcal{N}(0, v_1) + (1 - \alpha) \mathcal{N}(0, v_2),$$

is the marginal of

$$\begin{aligned} p(e_i, l_i | v_1, v_2) &= p(e_i | l_i, v_1, v_2) p(l_i) \\ p(e_i | l_i = \epsilon_1, v_1) &= \mathcal{N}(0, v_1), \\ p(e_i | l_i = \epsilon_2, v_2) &= \mathcal{N}(0, v_2), \end{aligned}$$

Student-t

$p(e_i) = \mathcal{St}(\nu, 0, \sigma)$, is the marginal of

$$\begin{aligned} p(e_i) &= \int p(e_i | \beta_i) p(\beta_i) \\ p(e_i | \beta_i) &= \mathcal{N}(0, \beta_i), \\ p(\beta_i) &= i\mathcal{G}(\gamma_0, \delta_0), \end{aligned}$$

Difference between S&S and Student

- ▶ minimum and maximum variance?
- ▶ the number of “clusters”?

Spike&Slab

Spike and slab prior of the noise implies:

$$\begin{aligned}p(e_i, l_i) &= \mathcal{N}(0, v_1)^{l_{1,i}} \mathcal{N}(0, v_2)^{l_{2,i}}, \\p(y_i, l_i | x_i) &= \mathcal{N}(ax_i + b, v_1)^{l_{1,i}} \mathcal{N}(ax_i + b, v_2)^{l_{2,i}},\end{aligned}$$

Do the point outside of the line follow the $ax_i + b$ curve?

Spike&Slab

Spike and slab prior of the noise implies:

$$\begin{aligned}p(e_i, l_i) &= \mathcal{N}(0, v_1)^{l_i} \mathcal{N}(0, v_2)^{1-l_i}, \\p(y_i, l_i | x_i) &= \mathcal{N}(ax_i + b, v_1)^{l_i} \mathcal{N}(ax_i + b, v_2)^{1-l_i},\end{aligned}$$

Do the point outside of the line follow the $ax_i + b$ curve?

What model do they follow?

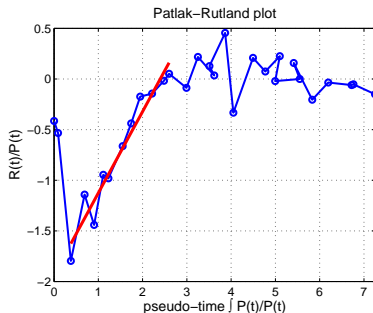
$$\begin{aligned}p(y_i, l_i | x_i) &= \mathcal{N}(ax_i + b, v_1)^{l_i} \\p(x_i | l_i = \epsilon_2) &= \epsilon_2^{l_i},\end{aligned}$$

Possibilities:

$$\begin{aligned}p(x_i | l_i = \epsilon_2) &= \mathcal{U}(y_{min}, y_{max}), \\p(x_i | l_i = \epsilon_2) &= \mathcal{N}(ax, v(x_i, \psi)),\end{aligned}$$

⋮

Be creative.



Tutorial model

Recall matrix notation from linear regression:

$$y_i = \mathbf{x}_i^\top \theta, \quad y_i = [\mathbf{x}_i \quad 1] [a \quad b]^\top,$$

With one-hot label $l_i = [1, 0]$ if in regression $l_i = [0, 1]$ if outside

$$p(y_i | x_i, l_i) = \mathcal{N}(\mathbf{x}_i^\top \theta, \omega^{-1})^{l_{1,i}} U(-2, 1)^{l_{2,i}},$$

$$p(l_i | \alpha) = \alpha_1^{l_{1,i}} \alpha_2^{l_{2,i}}, \quad p(\alpha) = \text{Di}(z_0)$$

$$p(\theta) = \mathcal{N}(0, \tau I_2), \quad p(\omega) = \Gamma(\gamma_0, \delta_0)$$

Joint distribution:

$$\begin{aligned} \mathcal{L} = \sum_{i=1}^n \log p(y_i, l_i | x_i, \theta, \omega, \alpha) p(\theta, \omega, \alpha) &= \frac{1}{2} \sum_i l_{1,i} (\log(\omega/2\pi) - \omega(y_i - \mathbf{x}_i^\top \theta)^2) \\ &+ \sum_i l_{2,i} \log \frac{1}{3} + \sum_i \sum_k l_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k + \\ &+ \frac{2}{2} \log \tau - \frac{1}{2} \tau \theta^\top \theta + (\gamma_0 - 1) \log \omega - \delta_0 \omega, \\ \text{s.t. } \sum_k l_{i,k} &= 1, \quad \sum_k \alpha_k = 1, \end{aligned}$$

Categorical distribution

Categorical distribution $i \in \{c_1, c_2 \dots c_K\}$

$$p(i|\alpha) = \prod_{k=1}^K \alpha_k^{[i=c_k]}$$

Categorical distribution

Categorical distribution $i \in \{c_1, c_2 \dots c_K\}$

$$p(i|\alpha) = \prod_{k=1}^K \alpha_k^{[i=c_k]}$$
$$\sum \alpha_k = 1$$

where $[i = c_k]$ is the Iverson bracket.

With one-hot $l \in \{[1, 0 \dots 0], [0, 1, \dots 0] \dots [0, 0 \dots, 1]\}$

$$p(l|\alpha) = \prod_k \alpha_k^{l_k}$$
$$\log p(l|\alpha) = \sum_k l_k \log \alpha_k \quad E(l) = \alpha, \hat{l} = \alpha$$

In mixtures,

$$\log p(l|\dots) \propto \sum_k l_k \mathcal{E}_k(\dots),$$

Categorical distribution

Categorical distribution $i \in \{c_1, c_2 \dots c_K\}$

$$p(i|\alpha) = \prod_{k=1}^K \alpha_k^{[i=c_k]}$$
$$\sum \alpha_k = 1$$

where $[i = c_k]$ is the Iverson bracket.

With one-hot $l \in \{[1, 0 \dots 0], [0, 1, \dots 0] \dots [0, 0 \dots, 1]\}$

$$p(l|\alpha) = \prod_k \alpha_k^{l_k}$$
$$\log p(l|\alpha) = \sum_k l_k \log \alpha_k \quad E(l) = \alpha, \hat{l} = \alpha$$

In mixtures,

$$\log p(l|\dots) \propto \sum_k l_k \mathcal{E}_k(\dots),$$
$$\hat{l}_k = \frac{\exp \mathcal{E}_k(\dots)}{\sum \exp \mathcal{E}_k(\dots)}$$

Expectation maximization (EM) algorithm: M-step

Expectation over l : $l_{1,i} \rightarrow \hat{l}_{1,i}$, $l_{2,i} \rightarrow \hat{l}_{2,i}$

$$\begin{aligned} E_l(\mathcal{L}) &= \frac{1}{2} \sum_i \hat{l}_{1,i} \log \omega - \frac{1}{2} \sum_i \hat{l}_{1,i} \omega (y_i - \mathbf{x}_i^\top \theta)^2 \\ &\quad + \sum_i \hat{l}_{2,i} \log \frac{1}{3} + \sum_i \sum_k \hat{l}_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k + \\ &\quad + \frac{2}{2} \log \tau - \frac{1}{2} \tau \theta^\top \theta + (\gamma_0 - 1) \log \omega - \delta_0 \omega + \text{n.c.}, \quad \text{s.t. : } \sum_k \alpha_k = 1, \end{aligned}$$

In θ (via moment matching of $q(\theta) = \mathcal{N}(\hat{\theta}, \Sigma_\theta)$):

$$\begin{aligned} \mathcal{L} &= -\frac{1}{2} \sum_i \hat{l}_{1,i} \omega (y_i - \mathbf{x}_i^\top \theta)^2 - \frac{1}{2} \tau \theta^\top \theta \\ &= -\frac{1}{2} \sum_i \hat{l}_{1,i} \omega (y_i^2 - 2y_i \mathbf{x}_i^\top \theta + \theta^\top \mathbf{x}_i \mathbf{x}_i^\top \theta) - \frac{1}{2} \tau \theta^\top \theta \\ &= -\frac{1}{2} \left(\omega \sum_i \hat{l}_{1,i} y_i^2 - 2\omega \sum_i \hat{l}_{1,i} y_i \mathbf{x}_i^\top \theta + \theta^\top \left(\omega \sum_i \hat{l}_{1,i} \mathbf{x}_i \mathbf{x}_i^\top + \tau l_2 \right) \theta \right) \\ \hat{\theta} &= \left(\omega \sum_i \hat{l}_{1,i} \mathbf{x}_i \mathbf{x}_i^\top + \tau l_2 \right)^{-1} \left(\omega \sum_i \hat{l}_{1,i} y_i \mathbf{x}_i^\top \right) \end{aligned}$$

Expectation maximization (EM) algorithm: M-step

Expectation over l : $l_{1,i} \rightarrow \hat{l}_{1,i}$, $l_{2,i} \rightarrow \hat{l}_{2,i}$

$$\begin{aligned}\mathcal{L} = & \frac{1}{2} \sum_i \hat{l}_{1,i} \log \omega - \frac{1}{2} \sum_i \hat{l}_{1,i} \omega (y_i - \mathbf{x}_i^\top \theta)^2 \\ & + \sum_i \hat{l}_{2,i} \log \frac{1}{3} + \sum_i \sum_k \hat{l}_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k + \\ & + \frac{2}{2} \log \tau - \frac{1}{2} \tau \theta^\top \theta + (\gamma_0 - 1) \log \omega - \delta_0 \omega + \text{n.c.}, \quad \text{s.t. : } \sum_k \alpha_k = 1,\end{aligned}$$

In ω , (via moment matching of $q(\omega) = \mathcal{G}(\gamma, \delta)$):

$$\begin{aligned}\mathcal{L} &= (\gamma - 1) \log \omega - \delta \omega \\ \gamma &= \frac{1}{2} \sum_i \hat{l}_{1,i} + \gamma_0 \\ \delta &= \frac{1}{2} \sum_i \hat{l}_{1,i} (y_i - \mathbf{x}_i^\top \hat{\theta})^2 + \delta_0 \\ \hat{\omega} &= \gamma / \delta\end{aligned}$$

Expectation maximization (EM) algorithm: M-step

Expectation over l : $l_{1,i} \rightarrow \hat{l}_{1,i}$, $l_{2,i} \rightarrow \hat{l}_{2,i}$

$$\begin{aligned}\mathcal{L} = & \frac{1}{2} \sum_i \hat{l}_{1,i} \log \omega - \frac{1}{2} \sum_i \hat{l}_{1,i} \omega (y_i - \mathbf{x}_i^\top \theta)^2 \\ & + \sum_i \hat{l}_{2,i} \log \frac{1}{3} + \sum_i \sum_k \hat{l}_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k + \\ & + \frac{2}{2} \log \tau - \frac{1}{2} \tau \theta^\top \theta + (\gamma_0 - 1) \log \omega - \delta_0 \omega + \text{n.c.}, \quad \text{s.t. : } \sum_k \alpha_k = 1,\end{aligned}$$

In α (via moment matching of $q(\alpha) = \mathcal{D}i(z)$):

$$\begin{aligned}\mathcal{L} = & \sum_k \sum_i \hat{l}_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k, \quad \text{s.t. : } \sum_k \alpha_k = 1, \\ = & \sum_k z_k \log \alpha_k + \lambda (\sum_k \alpha_k - 1) \\ z_k = & \sum_i \hat{l}_{k,i} + z_0 \\ \hat{\alpha}_k = & z_k / \sum_k z_k\end{aligned}$$

Expectation maximization (EM) algorithm: E-step

Probability of l_i given estimates

$$\begin{aligned}\mathcal{L}(l_i) = & \frac{1}{2} l_{1,i} \log \omega - \frac{1}{2} l_{1,i} \omega (y_i - \mathbf{x}_i^\top \theta)^2 - l_{1,i} \frac{1}{2} \log(2\pi) \\ & + l_{2,i} \log \frac{1}{3} + \sum_k l_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k + \text{n.c.}, \quad \text{s.t. : } \sum_k l_k = 1\end{aligned}$$

Template:

$$p(l_i|w) \propto \prod_k w_k^{l_{k,i}}, \quad \sum_k w_k = 1, \quad \sum_k l_{k,i} = 1 \text{ (in our case)}$$

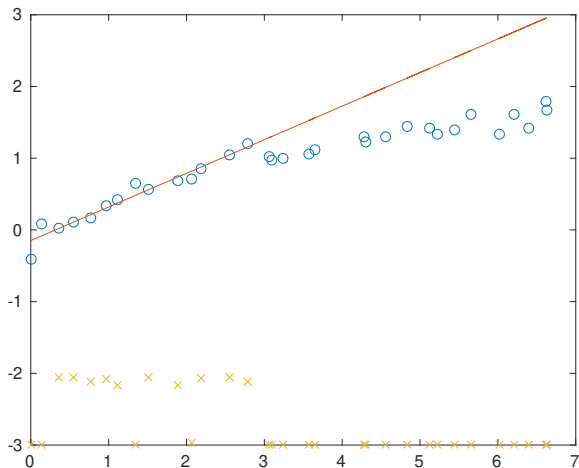
Matching (check with the Bayes rule):

$$\log w_1 = \frac{1}{2} \log \omega - \frac{1}{2} \omega (y_i - \mathbf{x}_i^\top \theta)^2 - \frac{1}{2} \log(2\pi) + \log \alpha_1$$

$$\log w_2 = \log \frac{1}{3} + \log \alpha_2$$

$$\hat{l}_1 = w_1 / (w_1 + w_2)$$

EM algorithm results



► sensitive to initialization

Modeling choices

How to encode our knowledge that the linear part is somewhere between 1 and 3min? hard constraints should be avoided.

Modeling choices

How to encode our knowledge that the linear part is somewhere between 1 and 3min? hard constraints should be avoided.

1. Prior on the latent l :

$$p(l|x) = \alpha_0(x),$$

2. Other options? Variance of the noise?

Assignment

Load data Patlak.mat

35 studies with:

`xpr` x axis

`ypr` y axis

`name` name of the study

`int_start` ignore

`int_end` ignore

Assignment	points
find slope of linear part for all 35 studies	
own code (WLS + uniform)	5
own code (WLS + novel proposition)	10