

# Mixture Model Challenge

Václav Šmíd

March 31, 2020

# Patlak Rutland plot

## Sequence of scintigraphic images of kidneys.

filename: drsprg\_023

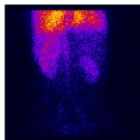
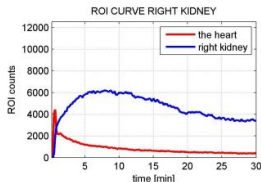
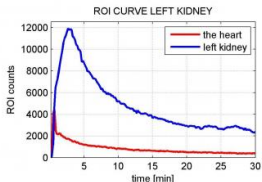
gender = F, age = 31 yrs

CKD stage = 2, LK = 77 %

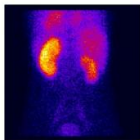
serum Cr - 0, Cr clearance - 0

99mTc-MAG3 - 0, 51Cr-EDTA - 0

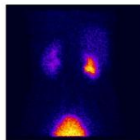
57Co-FLOOD - 1



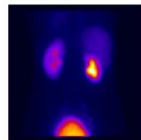
0 - 1 min



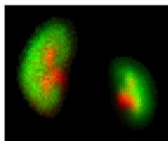
1 - 2 min



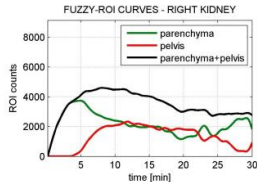
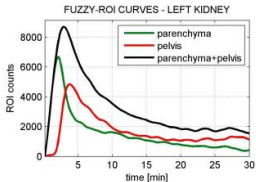
29 - 30 min



MEAN IMAGE



FUZZY ROIS

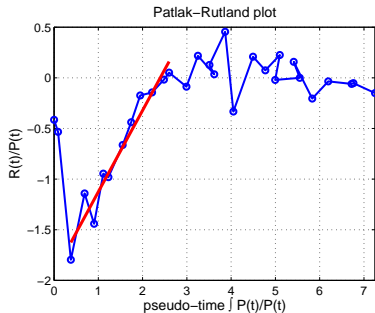


# Challenge: Patlak Rutland plot

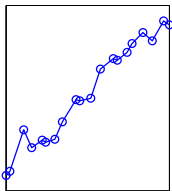
Patlak Rutland plot is a ratio of the parenchyma curve over the cumulative sum of the heart curve.

Diagnostically important is the **slope** of the linear part, which:

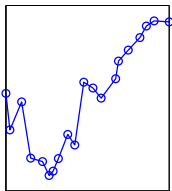
- ▶ typically starts around 1min
- ▶ typically ends around 3min
- ▶ with outliers
- ▶ manual fit is the norm



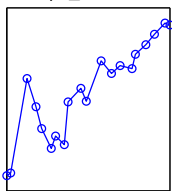
dtpa\_1i1c.crv



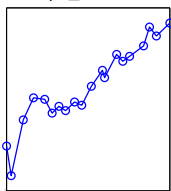
dtpa\_2i1c.crv



dtpa\_3i1c.crv



dtpa\_4i1c.crv

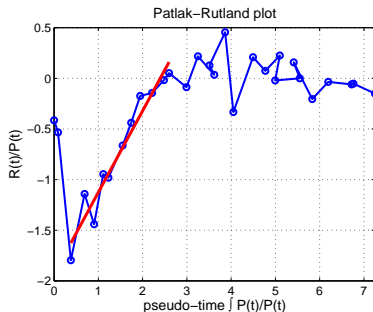


# Models?

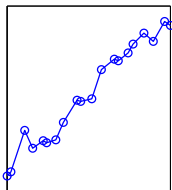
- ▶ no heuristics
  - ▶ probabilistic model providing either:
    - ▶ estimate of  $\hat{a}$
    - ▶ posterior distribution  $p(a)$
- where  $a$  is the slope of:

$$y_i = ax_i + b + e_i$$

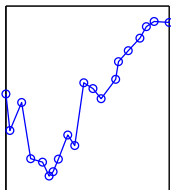
for points belonging to it.



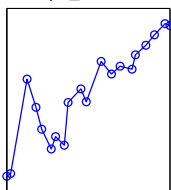
dtpa\_1i1c.crv



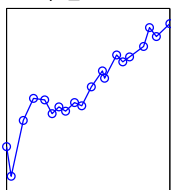
dtpa\_2i1c.crv



dtpa\_3i1c.crv



dtpa\_4i1c.crv

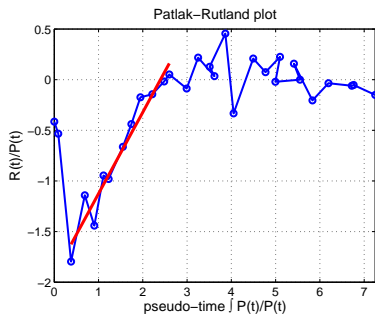


# Model 1: Robust regression

Model

$$y_i = ax_i + b + e_i$$

is valid for all  $x_i \in [1, 3]$  with heavy-tailed noise  $p(e_i)$ .



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## Model

$$y_i = ax_i + b + e_i$$

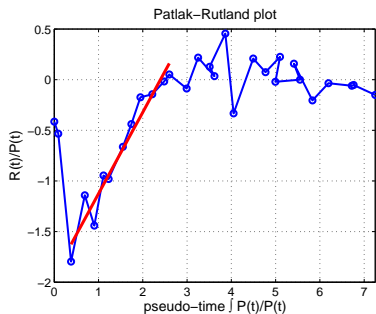
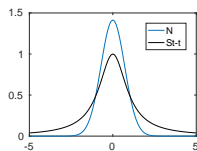
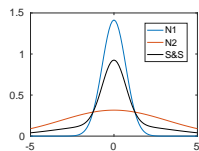
is valid for all  $x_i \in [1, 3]$  with heavy-tailed noise  $p(e_i)$ .

Possibilities:

- ▶ Spike and slab,

$$p(e_i) = \alpha \mathcal{N}(0, v_1) + (1-\alpha) \mathcal{N}(0, v_2),$$

- ▶ Student-t  $p(e_i) = \mathcal{St}(\nu, 0, \sigma)$ ,



# Latent representations

## Spike and slab

$$p(e_i | v_1, v_2) = \alpha \mathcal{N}(0, v_1) + (1 - \alpha) \mathcal{N}(0, v_2),$$

is the marginal of

$$\begin{aligned} p(e_i, l_i | v_1, v_2) &= p(e_i | l_i, v_1, v_2) \\ p(e_i | l_i = \epsilon_1, v_1) &= \mathcal{N}(0, v_1), \\ p(e_i | l_i = \epsilon_2, v_2) &= \mathcal{N}(0, v_2), \end{aligned}$$

## Student-t

$p(e_i) = \mathcal{St}(\nu, 0, \sigma)$ , is the marginal of

$$\begin{aligned} p(e_i) &= \int p(e_i | \beta_i) p(\beta_i) \\ p(e_i | \beta_i) &= \mathcal{N}(0, \beta_i), \\ p(\beta_i) &= i\mathcal{G}(\gamma_0, \delta_0), \end{aligned}$$

# Spike&Slab

Spike and slab prior of the noise implies:

$$p(e_i, l_i) = \mathcal{N}(0, v_1)^{l_i=\epsilon_1} \mathcal{N}(0, v_2)^{l_i=\epsilon_2},$$
$$p(y_i, l_i | x_i) = \mathcal{N}(ax_i + b, v_1)^{l_i=\epsilon_1} \mathcal{N}(ax_i + b, v_2)^{l_i=\epsilon_2},$$

Do the point outside of the line follow the  $ax_i + b$  curve?



# Spike&Slab

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$$p(y_i, l_i | x_i) = \mathcal{N}(ax_i + b, v_1)^{l_i=\epsilon_1} \mathcal{N}(ax_i + b, v_2)^{l_i=\epsilon_2},$$

Do the point outside of the line follow the  $ax_i + b$  curve?

What model do they follow?

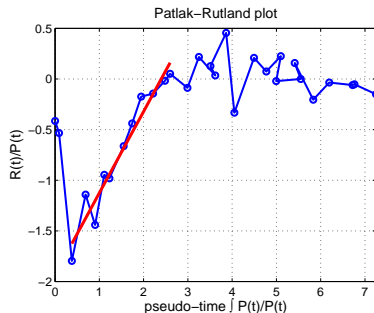
$$p(y_i, l_i | x_i) = \mathcal{N}(ax_i + b, v_1)^{l_i=\epsilon_1}$$
$$p(x_i | l_i = \epsilon_2)^{l_i=\epsilon_2},$$

Possibilities:

$$p(x_i | l_i = \epsilon_2) = \mathcal{U}(y_{min}, y_{max}),$$
$$p(x_i | l_i = \epsilon_2) = \mathcal{N}(ax, v(x_i, \psi)),$$

⋮

Be creative.



# Tutorial model

Recall matrix notation from linear regression:

$$y_i = \mathbf{x}_i^\top \theta, \quad y_i = [x_i \quad 1] [a \quad b]^\top,$$

Then

$$p(y_i|x_i, l_i) = \mathcal{N}(\mathbf{x}_i^\top \theta, \sigma)^{l_{1,i}} U(-2, 1)^{l_{2,i}},$$

$$p(l_i|\alpha) = \prod_{k=1}^K \alpha_k^{l_{k,i}} \quad p(\alpha) = \prod_k \alpha_k^{z_0-1}$$

$$p(\theta) = \mathcal{N}(0, \tau l_2), \quad p(\omega) = \Gamma(\gamma_0, \delta_0)$$

Joint distribution:

$$\mathcal{L} = \sum_{i=1}^n \log p(y_i, l_i|x_i, \theta, \omega, \alpha) p(\theta, \omega, \alpha) = \frac{1}{2} \sum_i l_{1,i} (\log(\omega/2\pi) - \omega(y_i - \mathbf{x}_i^\top \theta)^2)$$

$$+ \sum_i l_{2,i} \log \frac{1}{3} + \sum_i \sum_k l_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k +$$

$$+ \frac{2}{2} \log \tau - \frac{1}{2} \tau \theta^\top \theta + (\gamma - 1) \log \omega - \delta_0 \omega,$$

$$\text{s.t. } \sum l_{i,k} = 1, \quad \sum \alpha_k = 1,$$

# Expectation maximization (EM) algorithm: M-step

Expectation over  $l$ :  $l_{1,i} \rightarrow \hat{l}_{1,i}$ ,  $l_{2,i} \rightarrow \hat{l}_{2,i}$

$$\begin{aligned} E_l(\mathcal{L}) &= \frac{1}{2} \sum_i \hat{l}_{1,i} \log \omega - \frac{1}{2} \sum_i \hat{l}_{1,i} \omega (y_i - \mathbf{x}_i^\top \theta)^2 \\ &\quad + \sum_i \hat{l}_{2,i} \log \frac{1}{3} + \sum_i \sum_k \hat{l}_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k + \\ &\quad + \frac{2}{2} \log \tau - \frac{1}{2} \tau \theta^\top \theta + (\gamma_0 - 1) \log \omega - \delta_0 \omega + \text{n.c.}, \quad \text{s.t. : } \sum_k \alpha_k = 1, \end{aligned}$$

In  $\theta$  (via moment matching of  $q(\theta) = \mathcal{N}(\hat{\theta}, \Sigma_\theta)$ ):

$$\begin{aligned} \mathcal{L} &= -\frac{1}{2} \sum_i \hat{l}_{1,i} \omega (y_i - \mathbf{x}_i^\top \theta)^2 - \frac{1}{2} \tau \theta^\top \theta \\ &= -\frac{1}{2} \sum_i \hat{l}_{1,i} \omega (y_i^2 - 2y_i \mathbf{x}_i^\top \theta + \theta^\top \mathbf{x}_i \mathbf{x}_i^\top \theta) - \frac{1}{2} \tau \theta^\top \theta \\ &= -\frac{1}{2} \left( \omega \sum_i \hat{l}_{1,i} y_i^2 - 2\omega \sum_i \hat{l}_{1,i} y_i \mathbf{x}_i^\top \theta + \theta^\top \left( \omega \sum_i \hat{l}_{1,i} \mathbf{x}_i \mathbf{x}_i^\top + \tau l_2 \right) \theta \right) \\ \hat{\theta} &= \left( \omega \sum_i \hat{l}_{1,i} \mathbf{x}_i \mathbf{x}_i^\top + \tau l_2 \right)^{-1} \left( \omega \sum_i \hat{l}_{1,i} y_i \mathbf{x}_i^\top \right) \end{aligned}$$

## Expectation maximization (EM) algorithm: M-step

Expectation over  $l$ :  $l_{1,i} \rightarrow \hat{l}_{1,i}$ ,  $l_{2,i} \rightarrow \hat{l}_{2,i}$

$$\begin{aligned}\mathcal{L} = & \frac{1}{2} \sum_i \hat{l}_{1,i} \log \omega - \frac{1}{2} \sum_i \hat{l}_{1,i} \omega (y_i - \mathbf{x}_i^\top \theta)^2 \\ & + \sum_i \hat{l}_{2,i} \log \frac{1}{3} + \sum_i \sum_k \hat{l}_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k + \\ & + \frac{2}{2} \log \tau - \frac{1}{2} \tau \theta^\top \theta + (\gamma_0 - 1) \log \omega - \delta_0 \omega + \text{n.c.}, \quad \text{s.t. : } \sum_k \alpha_k = 1,\end{aligned}$$

In  $\omega$ , (via moment matching of  $q(\omega) = \mathcal{G}(\gamma, \delta)$ ):

$$\begin{aligned}\mathcal{L} &= \gamma \log \omega - \delta \omega \\ \gamma &= \frac{1}{2} \sum_i \hat{l}_{1,i} + 1 + \gamma_0 - 1 \\ \delta &= \frac{1}{2} \sum_i \hat{l}_{1,i} (y_i - \mathbf{x}_i^\top \hat{\theta})^2 + \delta_0 \\ \hat{\omega} &= \gamma / \delta\end{aligned}$$

# Expectation maximization (EM) algorithm: M-step

Expectation over  $l$ :  $l_{1,i} \rightarrow \hat{l}_{1,i}$ ,  $l_{2,i} \rightarrow \hat{l}_{2,i}$

$$\begin{aligned}\mathcal{L} = & \frac{1}{2} \sum_i \hat{l}_{1,i} \log \omega - \frac{1}{2} \sum_i \hat{l}_{1,i} \omega (y_i - \mathbf{x}_i^\top \theta)^2 \\ & + \sum_i \hat{l}_{2,i} \log \frac{1}{3} + \sum_i \sum_k \hat{l}_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k + \\ & + \frac{2}{2} \log \tau - \frac{1}{2} \tau \theta^\top \theta + (\gamma_0 - 1) \log \omega - \delta_0 \omega + \text{n.c.}, \quad \text{s.t. : } \sum_k \alpha_k = 1,\end{aligned}$$

In  $\alpha$  (via moment matching of  $q(\alpha) = \mathcal{D}i(z)$ ):

$$\begin{aligned}\mathcal{L} = & \sum_k \sum_i \hat{l}_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k, \quad \text{s.t. : } \sum_k \alpha_k = 1, \\ = & \sum_k z_k \log \alpha_k + \lambda (\sum_k \alpha_k - 1) \\ z_k = & \sum_i \hat{l}_{k,i} + z_0 \\ \hat{\alpha}_k = & z_k / \sum_k z_k\end{aligned}$$

# Expectation maximization (EM) algorithm: E-step

Probability of  $l_i$  given estimates

$$\begin{aligned}\mathcal{L}(l_i) = & \frac{1}{2} l_{1,i} \log \omega - \frac{1}{2} l_{1,i} \omega (y_i - \mathbf{x}_i^\top \theta)^2 - l_{1,i} \frac{1}{2} \log(2\pi) \\ & + l_{2,i} \log \frac{1}{3} + \sum_k l_{k,i} \log \alpha_k + (z_0 - 1) \sum_k \log \alpha_k + \text{n.c.}, \quad \text{s.t. : } \sum_k l_k = 1\end{aligned}$$

Template:

$$p(l_i|w) \propto \prod_k w_k^{l_{k,i}}, \quad \sum_k w_k = 1, \quad \sum_k l_{k,i} = 1 (\text{in our case})$$

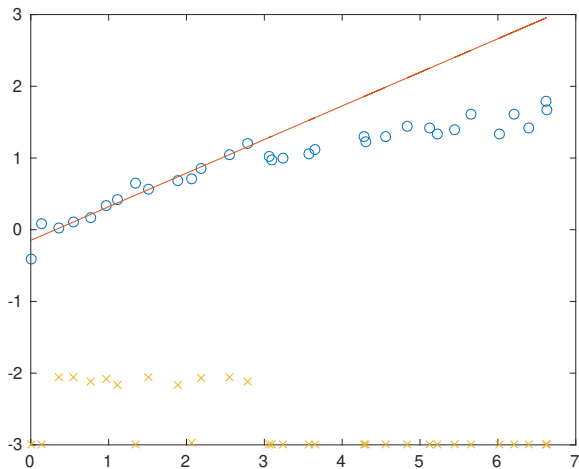
Matching (check with the Bayes rule):

$$\log w_1 = \frac{1}{2} \log \omega - \frac{1}{2} \omega (y_i - \mathbf{x}_i^\top \theta)^2 - \frac{1}{2} \log(2\pi) + \log \alpha_1$$

$$\log w_2 = \log \frac{1}{3} + \log \alpha_2$$

$$\hat{l}_1 = w_1 / (w_1 + w_2)$$

# EM algorithm results



► sensitive to initialization

# Modeling choices

How to encode our knowledge that the linear part is somewhere between 1 and 3min? hard constraints should be avoided.



# Modeling choices

How to encode our knowledge that the linear part is somewhere between 1 and 3min? hard constraints should be avoided.

1. Prior on the latent  $l$ :

$$p(l|x) = \alpha_0(x),$$

2. Other options? Variance of the noise?

# Assignment

Load data Patlak.mat

35 studies with:

`xpr` x axis

`ypr` y axis

`name` name of the study

`int_start` ignore

`int_end` ignore

Assignment	points
find slope of linear part for all 35 studies	
own code (WLS + uniform)	15
own code (WLS + novel proposition)	25