Interacting Particle Systems with Applications in Finance

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Lecture 1: Introduction

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- Interacting particle systems are mathematical models for collective behavior.
- Applications in physics (atoms & molecules), biology (organisms) & sociology, financial mathematics (people).
- Simple rules lead to complicated behavior.
- Markovian dynamics.
- Easy to simulate, but not always easy to prove; open problems.
- Rigorous methods lead to better understanding.

Let S be a finite set. A probability kernel on S is a function $K: S^2 \rightarrow [0, 1]$ such that $\sum_y K(x, y) = 1$. We calculate with kernels as with matrices:

$$\mathcal{KL}(x,z) := \sum_{y} \mathcal{K}(x,y)\mathcal{L}(y,z) \quad \text{and} \quad \mathcal{K}f(x) := \sum_{y} \mathcal{K}(x,y)f(y).$$

Let $X = (X_t)_{t \ge 0}$ be a stochastic process with values in S. By definition, X is a (time-homogeneous) *Markov process* if

$$\mathbb{P}\big[X_u \in \cdot \, \big| \, (X_s)_{0 \leq s \leq t}\big] = P_{u-t}(X_t, \, \cdot \,) \quad \text{a.s.} \qquad (0 \leq t \leq u),$$

where the *transition kernels* $(P_t)_{t\geq 0}$ form a collection of probability kernels on S such

$$P_sP_t = P_{s+t}$$
 and $\lim_{t\downarrow 0} P_t = P_0 = 1.$

Each such Markov semigroup $(P_t)_{t\geq 0}$ is of the form

$$P_t = e^{tG} := \sum_{n=0}^{\infty} \frac{1}{n!} (tG)^n,$$

where the generator G is a matrix of the form

$$G(x,y) \ge 0$$
 $(x \ne y)$ and $\sum_{y} G(x,y) = 0.$

We interpret G(x, y) $(x \neq y)$ as the *rate* of transitions $x \mapsto y$. The process $X = (X_t)_{t\geq 0}$ arises as the limit of discrete-time Markov chains with transition kernel of the form

$$P_{\varepsilon}(x,y) = 1_{\{x = y\}} + \varepsilon G(x,y) + O(\varepsilon^2).$$

Let $(\Omega, \mathcal{F}, \mu)$ be a measurable space with σ -finite, nonatomic measure μ . Recall that a *Poisson point set* with *intensity* μ is a random subset $\omega \subset \Omega$ such that

 $|\omega \cap A|$ is Poisson distributed with mean $\mu(A)$

whenever $A\in \mathcal{F}$, $\mu(A)<\infty$, and

$$|\omega \cap A_1|, \dots, |\omega \cap A_n|$$
 are independent (1)

whenever A_1, \ldots, A_n are disjoint. Since μ is nonatomic, for each $\varepsilon > 0$ we can find a countable partition $\{A_i^{\varepsilon} : i \in I\}$ of Ω such that $\mu(A_i^{\varepsilon}) \leq \varepsilon \quad \forall i$. Then $\mathbb{P}[|\omega \cap A_i^{\varepsilon}| = 1] = \mu(A_i^{\varepsilon}) + O(\varepsilon^2)$ and $\mathbb{P}[|\omega \cap A_i^{\varepsilon}| \geq 2] = O(\varepsilon^2)$. (2) Any sequence of random sets satisfying (1) and (2) converges as $\varepsilon \downarrow 0$ to a Poisson point set with intensity μ_{\cdot} Each generator G has a random mapping representation

$$Gf(x) = \sum_{m \in \mathcal{M}} r_m \{f(m(x)) - f(x)\},\$$

where $(r_m)_{m \in \mathcal{M}}$ are nonnegative rates and \mathcal{M} is a collection of maps $m : S \to S$. Let ω be a Poisson point set on $\mathcal{M} \times \mathbb{R}$ with intensity

$$\mu(\{m\} \times A) = r_m \ell(A) \qquad (A \in \mathcal{B}(\mathbb{R})),$$

where $\mathcal{B}(\mathbb{R})$ is the Borel- σ -field on \mathbb{R} and ℓ denotes Lebesgue measure. We may order the elements of

$$\omega \cap \mathcal{M} \times (s, t] =: \omega_{s,t} = \{(m_1, t_1), \ldots, (m_n, t_n)\}$$

with $t_1 < \cdots < t_n$.

Define random maps $\mathbf{X}_{s,t}:S
ightarrow S$ $(s\leq t)$ by

$$\mathbf{X}_{s,t} := m_n \circ \cdots \circ m_1.$$

(Poisson construction of Markov processes) Define maps $(\mathbf{X}_{s,t})_{s \leq t}$ as above in terms of a Poisson point set ω . Let X_0 be an S-valued random variable, independent of ω . Then

$$X_t := \mathbf{X}_{0,t}(X_0) \qquad (t \ge 0)$$

is a Markov process with generator G.

Remark The sample paths of X are cadlag, i.e., right-continuous with left limits. We get left-continuous paths by defining

$$\mathbf{X}_{s,t-}$$
 in terms of $\omega_{s,t-} := \omega \cap \mathcal{M} imes (s,t).$

Example 1 Consider the Markov process with state space $\{0,1\}$ that jumps with the rates

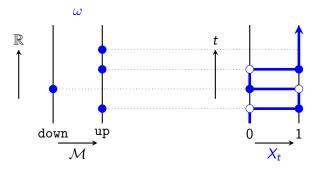
$$G(0,1) := 2$$
 and $G(1,0) := 1$.

Consider $\mathcal{M} := \{up, down\}$, where

$$up(x) := 1$$
 and $down(x) := 0$ $\forall x = 0, 1$.

Setting $r_{up} := 2$ and $r_{down} := 1$, we have

$$Gf(x) = r_{up} \{f(up(x)) - f(x)\} + r_{down} \{f(down(x)) - f(x)\}.$$



Since it may happen that $m(X_t) = X_t$, not every time of the Poisson process corresponds to a jump of the Markov process.

Random mapping representations are not unique!

Example 2 The same Markov process, that jumps

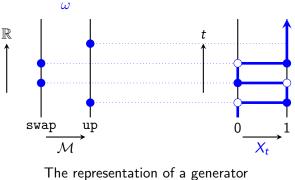
 $0\mapsto 1$ with rate 2 and $1\mapsto 0$ with rate 1

can also be represented as

$$Gf(x) = r_{up} \{f(up(x)) - f(x)\} + r_{swap} \{f(swap(x)) - f(x)\},\$$

with $r_{\mathrm{up}} := 1$, $r_{\mathrm{swap}} := 1$, and $\mathrm{up}(x) := 1$ and $\mathrm{swap}(x) := 1 - x$ (x = 0, 1).

Note that the total rate of jumps $0 \mapsto 1$ is $r_{up} + r_{swap} = 2$.



in terms of maps is not unique.

Let Λ be a finite set, representing people. Each person $i \in \Lambda$ can be in two states $x(i) \in \{-1, 1\}$. The state of the whole system is an element $x = (x(i))_{i \in \Lambda}$ of the space $S := \{-1, +1\}^{\Lambda}$ of all functions $x : \Lambda \to \{-1, +1\}$.

A person chooses his/her state according to an utility function. Given that the system is at time t in the state x, the utility for the person i of being in the state +1 resp. -1 is described by the utility function

$$U_t^{\pm}(i,x) = \pm \frac{1}{2}J\sum_{j\in\mathcal{N}_i}x(j) \pm \frac{1}{2}W_t(i).$$

Here $\mathcal{N}_i \subset \Lambda$ is a *neighborhood* of *i*, and $W_t(i)$ is a random, logistically distributed term.

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An Ising model for collective decision making

$$U_t^{\pm}(i,x) = \pm \frac{1}{2}J\sum_{j\in\mathcal{N}_i}x(j) \pm \frac{1}{2}W_t(i).$$

 For J > 0 it is advantageous to make the same choice as your neighbors.

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- $W_t(i)$ logistically distributed: $\mathbb{P}[W_t(i) \le w] = (1 + e^{-\beta w})^{-1}$.

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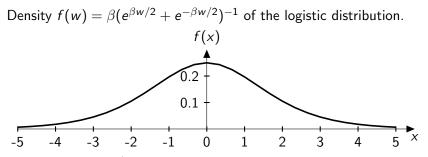
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- ► The noise terms W_t(i) are independent for each person and are redrawn at times of a Poisson process with intensity one.

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- ► The noise terms W_t(i) are independent for each person and are redrawn at times of a Poisson process with intensity one.
- After the noise is redrawn, each person immediately chooses his or her new state according to the highest utility.

An Ising model for collective decision making



Probability that $U_t^+(i, X_t) > U_t^-(i, X_t)$ given that \mathcal{N}_i contains $M_i(\pm)$ persons in the state ± 1 equals

 $\frac{e^{\beta JM_i(+)}}{e^{\beta JM_i(+)}+e^{\beta JM_i(-)}}.$

Person *i* changes his/her state to +1 with this rate and to -1 with one minus this rate. Only the product βJ matters.

Possible choices for the "neighborhood" \mathcal{N}_i of person *i*:

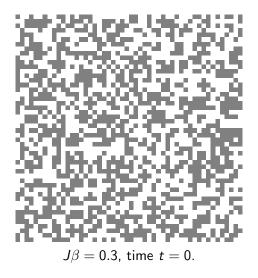
- 1. the whole set Λ ,
- 2. points $j \in \mathbb{Z}^d$ with |i j| = 1,
- 3. neighbors in any locally finite graph.

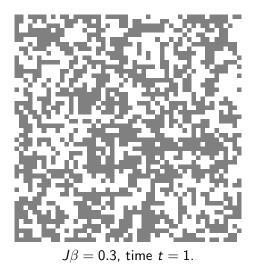
Choice 1. made by Brock & Durlauf (2001), who studied the invariant laws of this Markov chain (but not its time evolution).

Logistic distribution motivated by extreme-value theory, but mainly by the wish to obtain the *Ising model* (on \mathbb{Z}^d) or *Curie Weiss model* (on the complete graph), which are used in physics to describe systems of atoms at *inverse temperature* $1/\beta$ whose spin can be in two states ± 1 and that interact through magnetic forces.

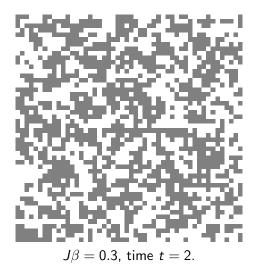
More precisely, our model is the Ising model with (continuous-time) *Glauber dynamics*.

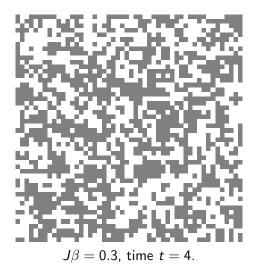
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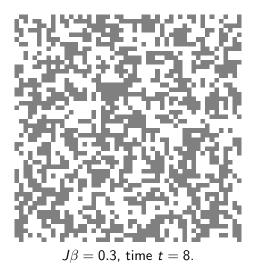


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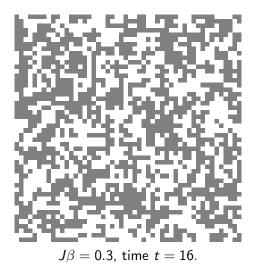


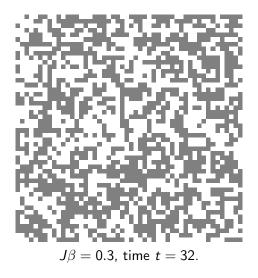


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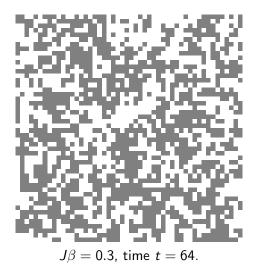


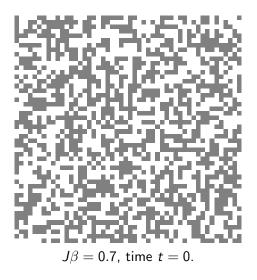
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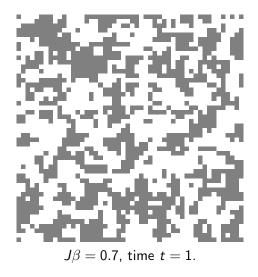


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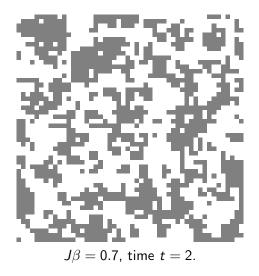


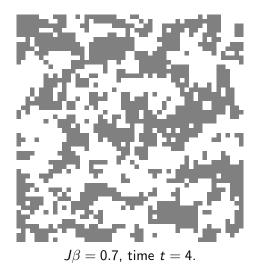


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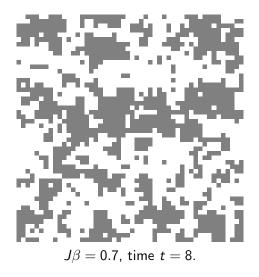


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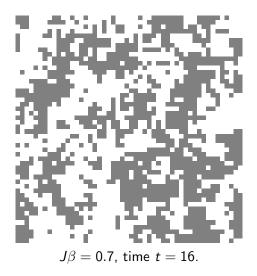




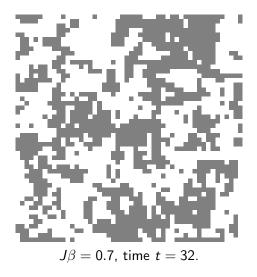
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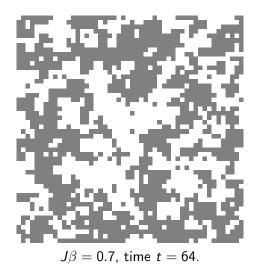
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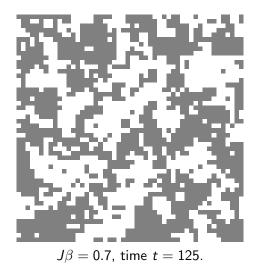
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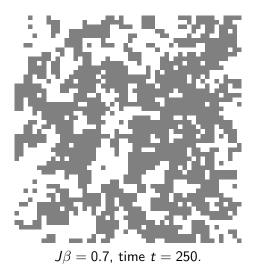
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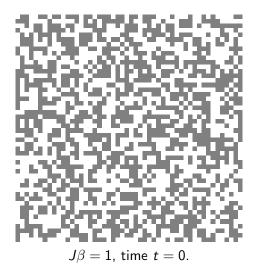
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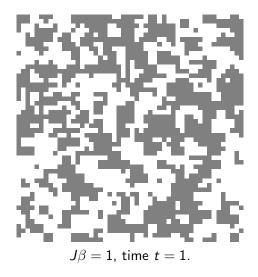
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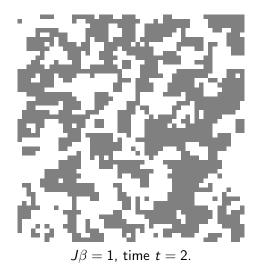
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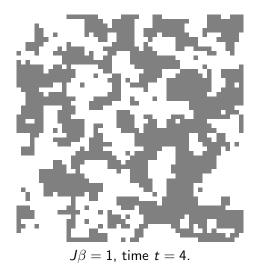
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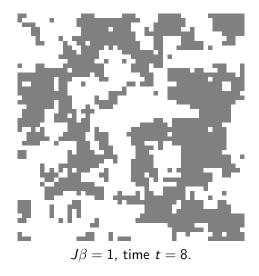
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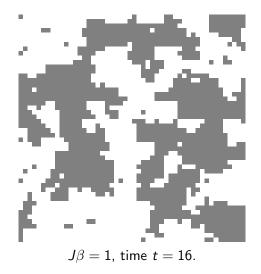


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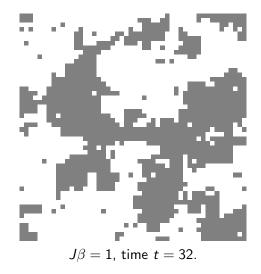
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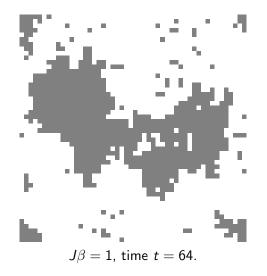
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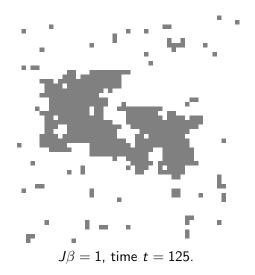


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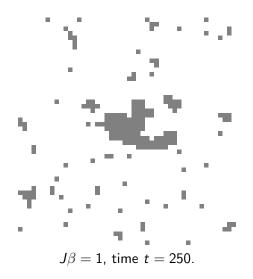


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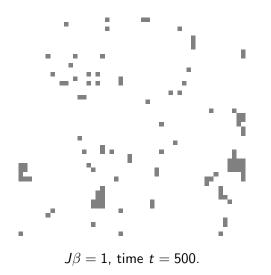


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Instead of allowing only two states -1, +1, we can more generally allow $q \ge 2$ states $1, \ldots, q$.

Each person i chooses a new state at times of a Poisson process with rate 1.

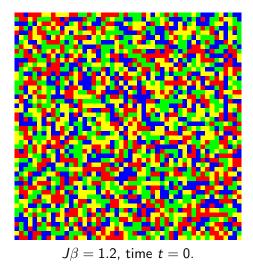
The probability that the newly chosen state is $k \in \{1, \dots, q\}$ equals

 $\frac{e^{\beta J M_i(k)}}{\sum_{m=1}^q e^{\beta J M_i(m)}},$

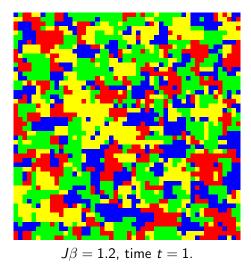
where $M_i(k)$ denotes the number of neighbors of *i* that are in the state *k*.

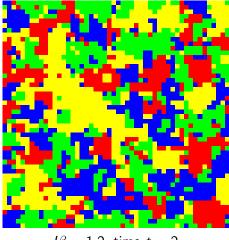
Special case q = 2 amounts to the Ising model.

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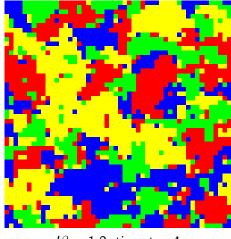


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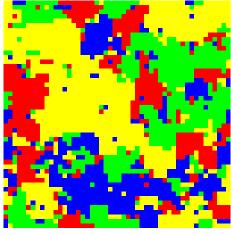


 $J\beta = 1.2$, time t = 2.



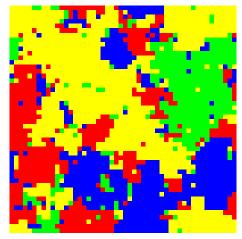
 $J\beta = 1.2$, time t = 4.

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 $J\beta = 1.2$, time t = 8.

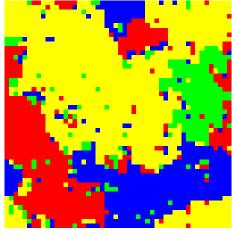
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 $J\beta = 1.2$, time t = 16.

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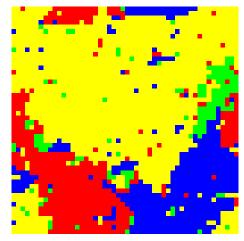
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 $J\beta = 1.2$, time t = 32.

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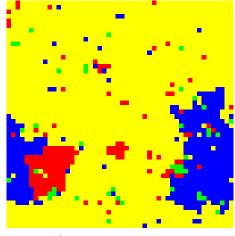
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 $J\beta = 1.2$, time t = 64.

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 $J\beta = 1.2$, time t = 125.

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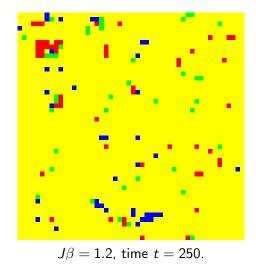
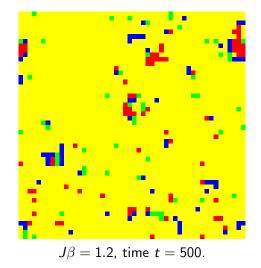


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In the *voter model*, persons can have $q \ge 2$ states.

Each person i chooses a new state at times of a Poisson process with rate 1.

The probability that the newly chosen state is $k \in \{1, \ldots, q\}$ equals

$$rac{1}{|\mathcal{N}_i|}M_i(k)$$

where $M_i(k)$ denotes the number of neighbors of *i* that are in the state *k*.

Contrary to the Potts model, types, once extinct, cannot reappear.

Used to model *credit contagion* in Gieseck & Weber (2002).

Also used to model *voting behavior*, or the spread of *neutral genetic types* in population biology.

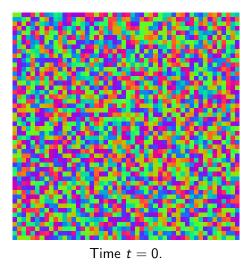
Using the voter map

$$\operatorname{vot}_{ij}(x) := \left\{ egin{array}{cc} x(j) & ext{if } k = i, \ x(k) & ext{otherwise,} \end{array}
ight.$$

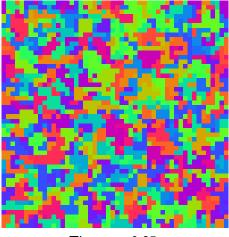
we can give the following random mapping representation of the generator:

$$Gf(x) = \sum_{i \in \Lambda} \frac{1}{|\mathcal{N}_i|} \sum_{j \in \mathcal{N}_i} \{f(\mathtt{vot}_{i,j}x) - f(x)\}.$$

Interpretation: each person copies with rate 1 the type of a uniformly chosen random person of N_i .

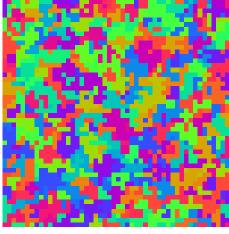


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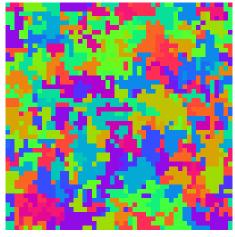


Time t = 0.25.

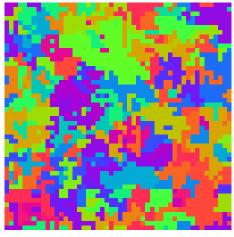
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Time t = 0.5.

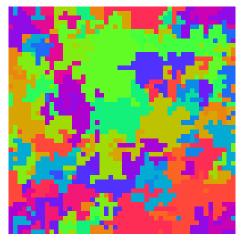


Time t = 1.



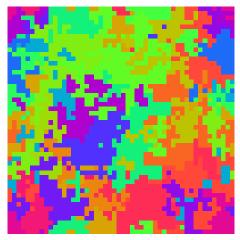
Time t = 2.

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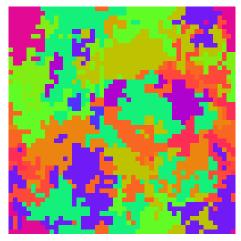
Time t = 4.

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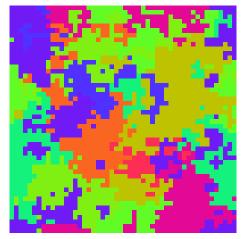
Time t = 8.

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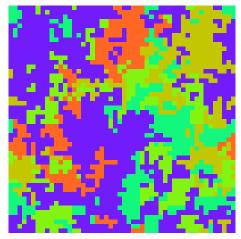
Time t = 16.

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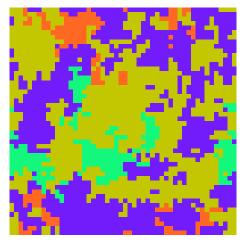
Time t = 31.25.

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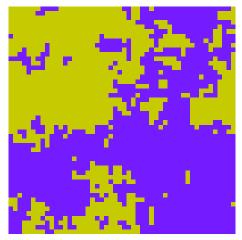
Time t = 62.5.

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Time t = 125.

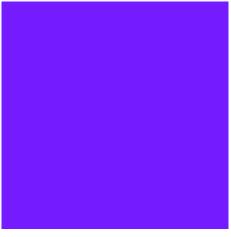
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Time t = 250.

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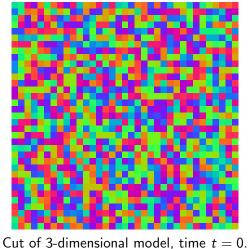
Time t = 500.

The behavior of the voter model strongly depends on the dimension.

Clustering in dimensions d = 1, 2.

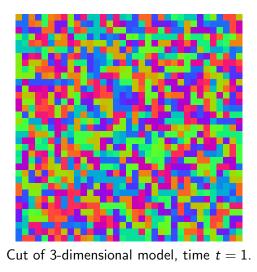
Stable behavior in dimensions $d \ge 3$.

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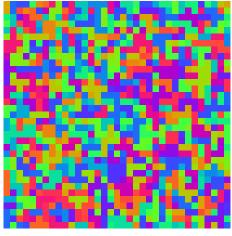
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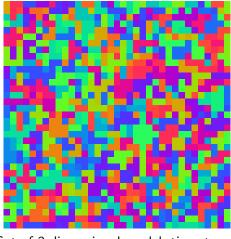
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Cut of 3-dimensional model, time t = 2.

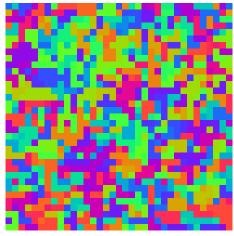
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Cut of 3-dimensional model, time t = 4.

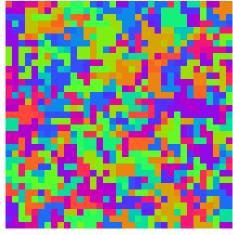
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Cut of 3-dimensional model, time t = 8.

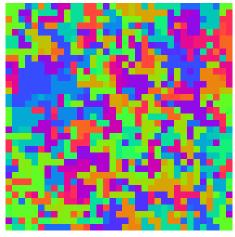
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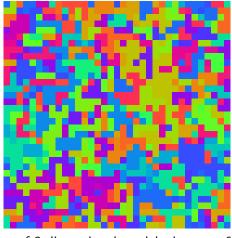
Cut of 3-dimensional model, time t = 16.

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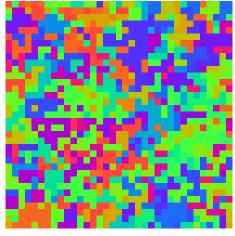
Cut of 3-dimensional model, time t = 32.

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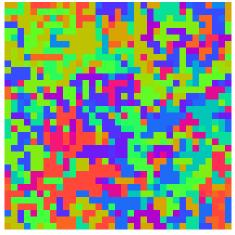
Cut of 3-dimensional model, time t = 64.

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Cut of 3-dimensional model, time t = 125.

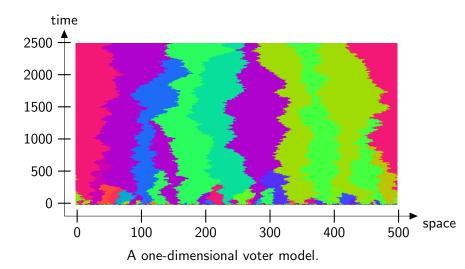
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Cut of 3-dimensional model, time t = 250.

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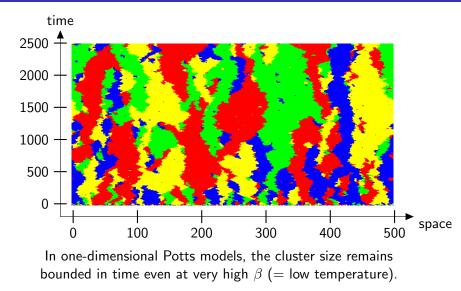
The voter model



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A one-dimensional Potts model



In the *biased* voter model with two states $\{0,1\}$, each person *i* changes its type $X_t(i)$ with the rates

- $0 \mapsto 1$ with rate $(1 + s) \cdot$ fraction of type 1 neighbors,
- $1\mapsto 0 \qquad \text{with rate } 1\cdot \text{fraction of type 0 neighbors},$

where s > 0 gives type 1 a (small) advantage.

Contrary to the voter model, even if we start with just a single person of type 1, there is a positive probability that type 1 never dies out.

Models spread of *new idea* or *technology*, or *advantageous mutation* in biology.

Biased voter model with s = 0.2. Time t = 0.

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Biased voter model with s = 0.2. Time t = 10.

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Biased voter model with s = 0.2. Time t = 20.

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Biased voter model with s = 0.2. Time t = 30.

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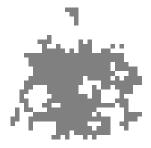
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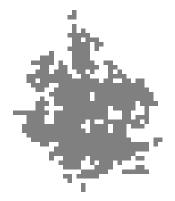


Biased voter model with s = 0.2. Time t = 40.

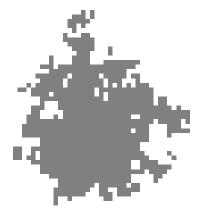
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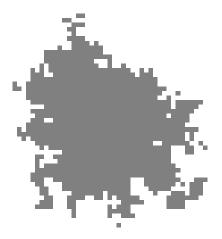
Biased voter model with s = 0.2. Time t = 50.



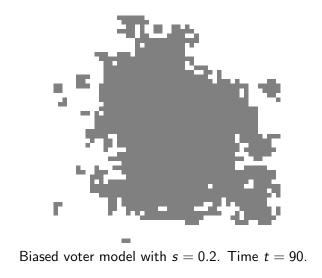
Biased voter model with s = 0.2. Time t = 60.

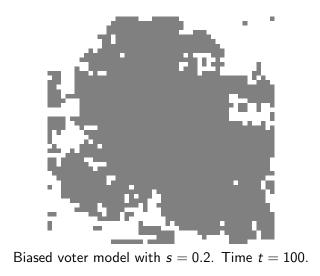


Biased voter model with s = 0.2. Time t = 70.



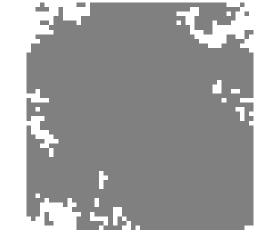
Biased voter model with s = 0.2. Time t = 80.







Biased voter model with s = 0.2. Time t = 110.



Biased voter model with s = 0.2. Time t = 120.

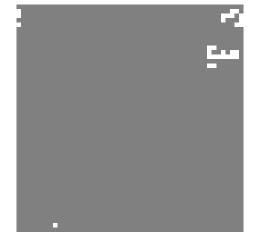


Biased voter model with s = 0.2. Time t = 130.

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Biased voter model with s = 0.2. Time t = 140.

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Biased voter model with s = 0.2. Time t = 150.

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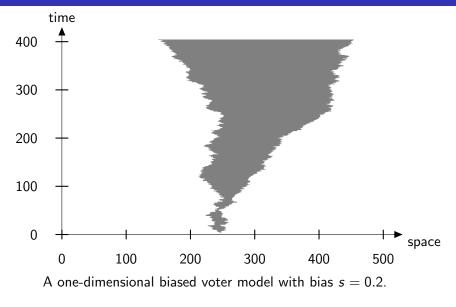
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Biased voter model with s = 0.2. Time t = 160.

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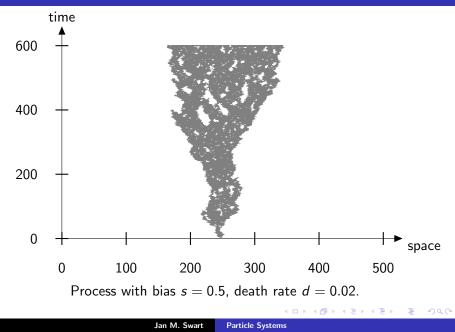
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We can extend the biased voter model by also allowing spontaneous jumps from 1 to 0.

$0\mapsto 1$	with rate $(1 + s) \cdot$ fraction of type 1 neighbors,
$1\mapsto 0$	with rate $1\cdot fraction$ of type 0 neighbors
	+ d,

where s > 0 gives type 1 an advantage and $d \ge 0$ is a *death rate*. This models the fact that complicated new ideas may be forgotten or organisms may die.

Whether 1's have a positive probability to survive now depends in a nontrivial way on s and d.



Simplifying, we can also look at a process that jumps as

- $0\mapsto 1 \qquad \text{with rate } \lambda \cdot \text{number of type 1 neighbors},$
- $1 \mapsto 0$ with rate d.

This is the *contact process* with *infection rate* $\lambda > 0$ and *death rate* d > 0.

Again, this can be used to model the spread of ideas or biological populations.

By changing the time scale, we can set the intensity of one Poisson process to one, i.e., without loss of generality d = 1 (the usual convention) or also $\lambda = 1$ (if we wish).

Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 0.

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Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 1.

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Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 2.

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Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 3.

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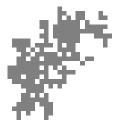
Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 4.

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Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 5.

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Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 6.

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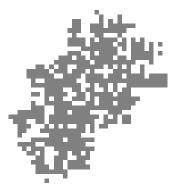


Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 7.

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Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 8.

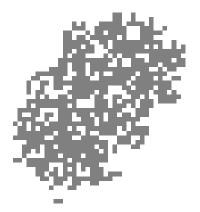


Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 9.

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Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 10.



Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 11.

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Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 12.



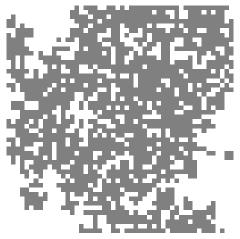


Time t = 14.

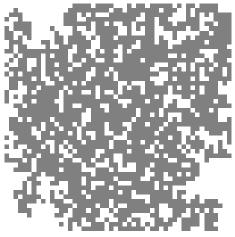


Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 15.

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Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 16.



Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 17.

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Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 18.

(4月) (日)



Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 19.

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Contact process with infection rate $\lambda = 2$ and death rate d = 1. Time t = 20.

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A rebellious voter model

Consider a model with two types $\{0,1\}$ and let

$$f_{ au} := rac{1}{|\mathcal{N}_i|} \sum_{j \in \mathcal{N}_i} \mathbb{1}\{x(j) = au\}$$

be the frequency of type τ in the neighborhood \mathcal{N}_i .

A person of type τ chooses a new type with rate

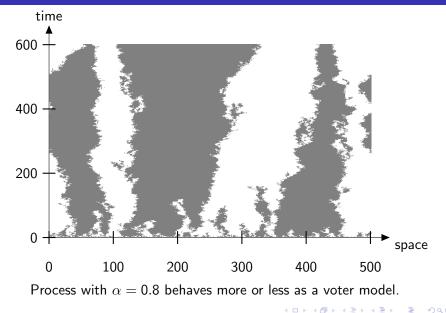
$$f_{\tau} + \alpha f_{1-\tau}.$$

For $\alpha < 1$, persons change their mind *less* often if they disagree with a lot of neighbors.

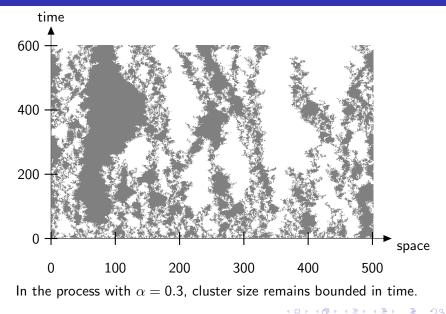
As in a normal voter model, the probability that the newly chosen type is τ' is $f_{\tau'}.$

Used by Neuhauser & Pacala (1999) to model balancing selection.

A rebellious voter model



A rebellious voter model



Reaction diffusion models

Another rich class of models are reaction diffusion models.

These are systems of particles that perform independent random walks and interact when they are near to each other.

Let $X_t(i) = 1$ (resp. 0) signify the presence (resp. absence) of a particle and consider the maps $rw_{ij} : \{0,1\}^{\mathbb{Z}} \to \{0,1\}^{\mathbb{Z}}$

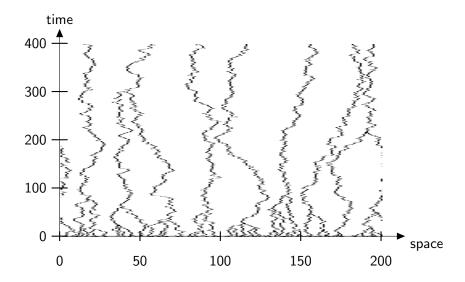
$$\mathbf{rw}_{i,j} \mathbf{x}(k) := \begin{cases} 0 & \text{if } k = i, \\ \mathbf{x}(i) \lor \mathbf{x}(j) & \text{if } k = j, \\ \mathbf{x}(k) & \text{otherwise.} \end{cases}$$

The process with generator

$$G = \frac{1}{2} \sum_{i \in \mathbb{Z}} \left\{ f(\texttt{rw}_{i,i+1}x) - f(x) \right\} + \frac{1}{2} \sum_{i \in \mathbb{Z}} \left\{ f(\texttt{rw}_{i,i-1}x) - f(x) \right\}$$

describes coalescing random walks.

Coalescing random walks



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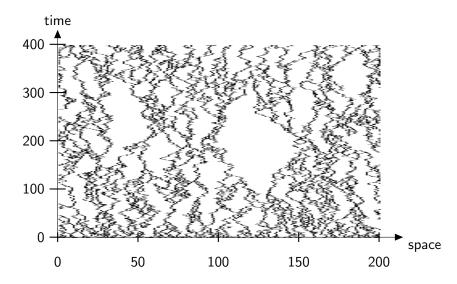
We can also add other maps to the dynamics, like the *branching map*

$$bra_{i,j}x(k) := \begin{cases} x(i) \lor x(j) & \text{if } k = j, \\ x(k) & \text{otherwise,} \end{cases}$$

or even cooperative branching

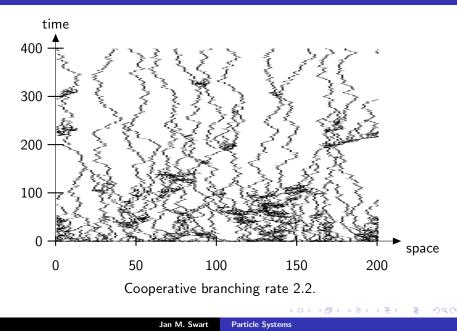
$$\operatorname{coop}_{i,i',j} x(k) := \begin{cases} (x(i) \wedge x(i')) \lor x(j) & \text{if } k = j, \\ x(k) & \text{otherwise.} \end{cases}$$

Branching and coalescing random walks

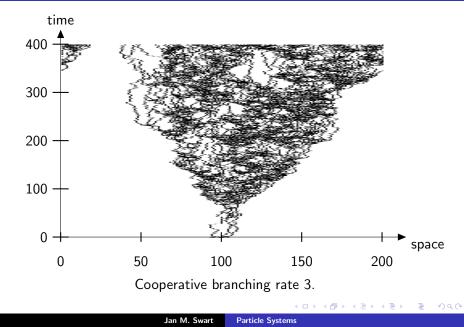


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Cooperative branching and coalescence



Cooperative branching



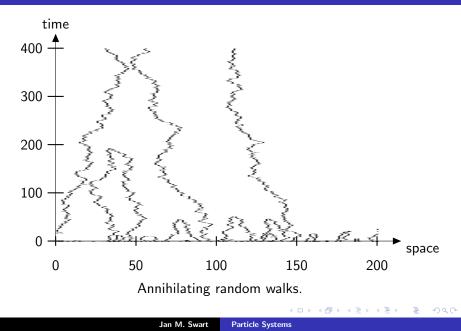
Two more maps of interest are the annihilating random walk map

$$\operatorname{arw}_{i,j} x(k) := \left\{ egin{array}{cc} 0 & ext{if } k=i, \ x(i)+x(j) \mod(2) & ext{if } k=j, \ x(k) & ext{otherwise}, \end{array}
ight.$$

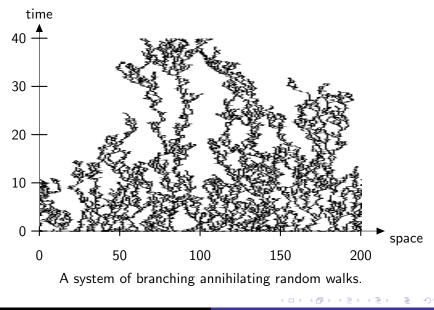
and the annihilating branching map

$$\mathtt{abra}_{i,j} x(k) := \left\{ egin{array}{ll} x(i) + x(j) \mod(2) & ext{if } k = j, \ x(k) & ext{otherwise}, \end{array}
ight.$$

A cancellative system



A cancellative system



Define a killing map as

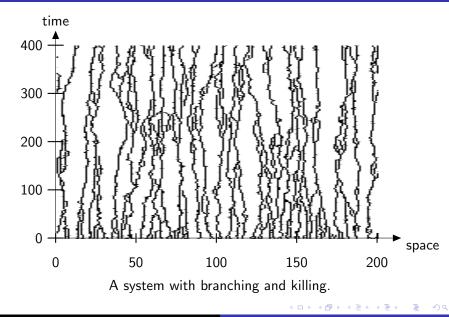
$$ext{kill}_{i,j} x(k) := \left\{egin{array}{cc} (1-x(i)) \wedge x(j) & ext{if } k=j, \ x(k) & ext{otherwise}, \end{array}
ight.$$

which says that the particle at i, if present, kills any particle at j.

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Branching and killing



Jan M. Swart Particle Systems