

Intertwining of Markov processes and the contact process on the hierarchical group

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Outline

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- ▶ First passage times of birth and death processes

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- ▶ The contact process on the hierarchical group

A bit of filtering theory

Let $(X, Y) = (X_n, Y_n)_{n \geq 0}$ be a Markov chain with finite state space of the form $S \times T$, which jumps from a point (x, y) to a point (x', y') with transition probability $P(x, y; x', y')$. Set

$$\mathbb{P}(y \mid x_0, \dots, x_n) := \mathbb{P}[Y_n = y \mid (X_0, \dots, X_n) = (x_0, \dots, x_n)].$$

The so-called *filtering equations* read:

$$\begin{aligned} \mathbb{P}(y \mid x_0, \dots, x_n) \\ = \frac{\sum_{y' \in S'} \mathbb{P}(y' \mid x_0, \dots, x_{n-1}) P(x_{n-1}, y'; x_n, y)}{\sum_{y', y'' \in S'} \mathbb{P}(y' \mid x_0, \dots, x_{n-1}) P(x_{n-1}, y'; x_n, y'')}. \end{aligned}$$

Markov functionals

Proposition Let K be a probability kernel from S to T . Assume that there exists a transition probability Q on S such that

$$Q(x, x')K(x', y') = \sum_y K(x, y)P(x, y; x', y')$$

Then

$$\mathbb{P}[Y_0 = y \mid X_0] = K(X_0, y) \quad \text{a.s.}$$

implies that

$$\mathbb{P}[Y_n = y \mid (X_0, \dots, X_n)] = K(X_n, y) \quad \text{a.s.} \quad (n \geq 0)$$

and X , on its own, is a Markov chain with transition kernel Q .

Continuous time

[Rogers & Pitman '81] Let K be a probability kernel from S to T and define $K : \mathbb{R}^{S \times T} \rightarrow \mathbb{R}^S$ by

$$Kf(x) := \sum_{y \in S'} K(x, y)f(x, y).$$

Let G and \hat{G} be Markov generators on S resp. $S \times T$. Assume that

$$GK = K\hat{G}.$$

Then the Markov process (X, Y) with generator \hat{G} started in an initial law such that

$$\mathbb{P}[Y_0 = y \mid X_0] = K(X_0, y) \quad \text{a.s.}$$

satisfies

$$\mathbb{P}[Y_t = y \mid (X_s)_{0 \leq s \leq t}] = K(X_t, y) \quad \text{a.s.} \quad (t \geq 0).$$

and X , on its own, is a Markov process with generator G .

Example: Wright-Fisher diffusion

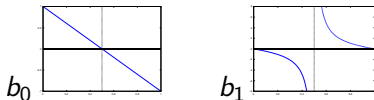
Let (X, Y) be the Markov process in $[0, 1] \times \{0, 1\}$ with generator

$$\hat{G}f(x, 0) = \frac{1}{2}x(1-x)\frac{\partial^2}{\partial x^2}f(x, 0) + b_0(x)\frac{\partial}{\partial x}f(x, 0) + (f(x, 1) - f(x, 0)),$$

$$\hat{G}f(x, 1) = \frac{1}{2}x(1-x)\frac{\partial^2}{\partial x^2}f(x, 1) + b_1(x)\frac{\partial}{\partial x}f(x, 1),$$

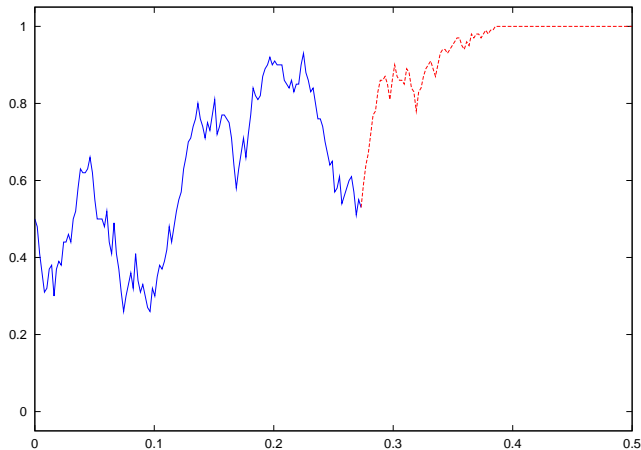
where

$$b_0(x) = 2\left(\frac{1}{2} - x\right), \quad b_1(x) = \frac{8x(1-x)(x - \frac{1}{2})}{1 - 4x(1-x)}.$$



Note The process Y is autonomous and jumps $0 \mapsto 1$ with rate one independent of the state of X , but the evolution of X depends on the state of Y .

Example: Wright-Fisher diffusion

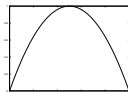


As long as
 $Y_t = 0$, X
cannot reach
the boundary.
When $Y_t = 1$,
 X cannot cross
the middle.

Example: Wright-Fisher diffusion

Define a kernel K from $[0, 1]$ to $\{0, 1\}$ by

$$K(x, 0) := 4x(1 - x), \quad K(x, 1) := 1 - 4x(1 - x)$$



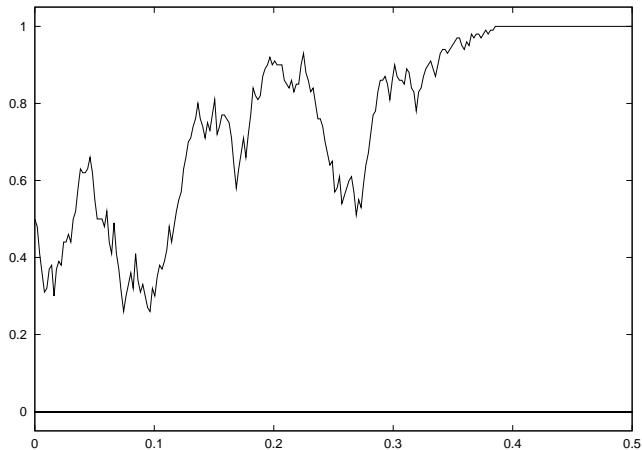
Then $(X_0, Y_0) = (\frac{1}{2}, 0)$ implies

$$\mathbb{P}[Y_t = y \mid (X_s)_{0 \leq s \leq t}] = K(X_t, y) \quad \text{a.s.} \quad (t \geq 0).$$

and X , on its own, is a Markov process with generator

$$Gf(x) = \frac{1}{2}x(1 - x) \frac{\partial^2}{\partial x^2} f(x).$$

Example: Wright-Fisher diffusion



Example: Wright-Fisher diffusion

Generalization It is possible to couple a Wright-Fisher diffusion X started in $X_0 = \frac{1}{2}$ to a process Y with state space $\{0, 1, \dots, \infty\}$, started in $Y_0 = 0$ such that Y jumps $k \mapsto k + 1$ with rate $(k + 1)(2k + 1)$, independent of the state of X , and:

$$\mathbb{P}[Y_t = k \mid (X_s)_{0 \leq s \leq t}] = (1 - X_t^2)X_t^{2k},$$

and

$$\inf \{t \geq 0 : X_t \in \{0, 1\}\} = \inf \{t \geq 0 : Y_t = \infty\}.$$

Adding structure to Markov processes

Let X be a Markov processes with state space S and generator G , let K be a probability kernel from S to T and let $(G'_x)_{x \in S}$ be a collection of generators of T -valued Markov processes. Assume that

$$GK = \overline{K} \overline{G}$$

where

$$\overline{K}f(x) := \sum_y K(x, y)f(x, y) \quad \text{and} \quad \overline{G}f(x, y) := G'_x f(y).$$

Then X can be coupled to a process Y such that (X, Y) is Markov, Y evolves according to the generator G'_x while X is in the state x , and

$$\mathbb{P}[Y_t = y \mid (X_s)_{0 \leq s \leq t}] = K(X_t, y) \quad \text{a.s.} \quad (t \geq 0).$$

We call Y an *added-on process on X* .

Intertwining of semigroups

Proposition Let X and Y be Markov processes with state spaces S and T , semigroups $(P_t)_{t \geq 0}$ and $(P'_t)_{t \geq 0}$, and generators G and G' . Let K be a probability kernel from S to T and define

$$Kf(x) := \sum_y K(x, y)f(y).$$

Then $GK = KG'$ implies the *intertwining relation*

$$P_t K = K P'_t \quad (t \geq 0)$$

and the processes X and Y can be coupled such that

$$\mathbb{P}[Y_t = y \mid (X_s)_{0 \leq s \leq t}] = K(X_t, y) \quad \text{a.s.} \quad (t \geq 0).$$

We call Y an *averaged Markov process on X* .

First passage times of birth and death processes

[Karlin & McGregor '59] Let Z be a Markov process with state space $\{0, 1, 2, \dots\}$, started in $Z_0 = 0$, that jumps $k - 1 \mapsto k$ with rate $b_k > 0$ and $k \mapsto k - 1$ with rate $d_k > 0$ ($k \geq 1$). Then

$$\tau_N := \inf\{t \geq 0 : Z_t = N\}$$

is distributed as a sum of independent exponentially distributed random variables whose parameters $\lambda_1 < \dots < \lambda_N$ are the negatives of the eigenvalues of the generator of the process stopped in N .

A probabilistic proof

[Diaconis & Miclos '09] Let Y be a Markov process with state space $\{0, 1, \dots, N\}$ that jumps $k-1 \mapsto k$ with rate λ_{N-k+1} . Then Y can be coupled to Z such that

$$\mathbb{P}[Z_{t \wedge \tau_N} = z \mid (Y_s)_{0 \leq s \leq t}] = K(Y_t, z) \quad \text{a.s.},$$

where K is a probability kernel on $\{0, \dots, N\}$ satisfying

$$K(N, N) = 1 \quad \text{and} \quad K(y, \{0, \dots, y\}) = 1 \quad (0 \leq y < N).$$

Note: here $Z_{t \wedge \tau_N}$ is an averaged Markov process on Y .

A probabilistic proof

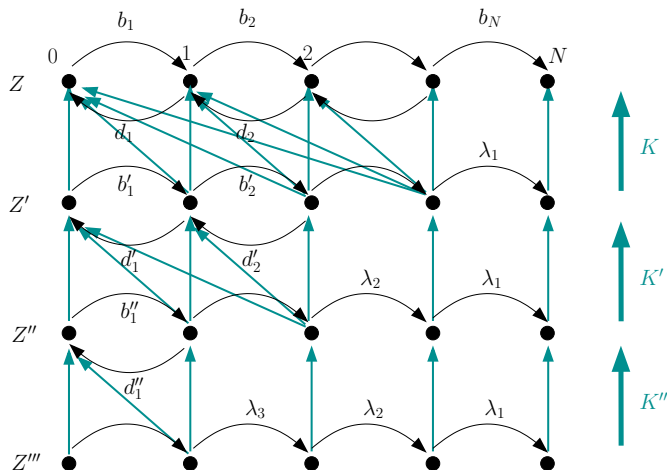
Proof By induction. Assume that Z is a process on $\{0, \dots, N\}$ such that $b_1, \dots, b_N > 0$, $d_1, \dots, d_{M-1} > 0$, $d_M = \dots = d_N = 0$ for some $1 \leq M \leq N$. Then Diaconis and Miclos construct a kernel K such that

$$\begin{aligned} K(y, \{0, \dots, y\}) &= 1 & (0 \leq y < M), \\ K(y, y) &= 1 & (M \leq y \leq N) \end{aligned}$$

and a process Z' with $b'_1, \dots, b'_N > 0$, $d'_1, \dots, d'_{M-2} > 0$, $d'_{M-1} = \dots = d'_N = 0$ such that

$$\mathbb{P}[Z_t = z \mid (Z'_s)_{0 \leq s \leq t}] = K(Z'_t, z) \quad \text{a.s.}$$

A probabilistic proof



A question

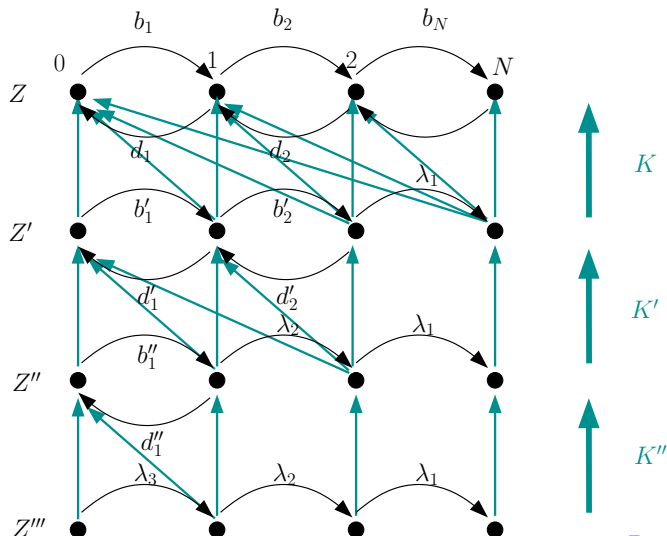
Question Diaconis and Miclos construct $Z_{t \wedge \tau_N}$ as an averaged Markov process on a process Y_t with zero death rates and with birth rates $\lambda_N, \dots, \lambda_1$, in such a way that $Z_{t \wedge \tau_N} \leq Y_t$. Let V_t be the process with zero death rates and with birth rates $\lambda_1, \dots, \lambda_N$ (in this order!). Is it possible to construct V_t as an averaged Markov process on $Z_{t \wedge \tau_N}$ (in this order!) in such a way that $V_t \leq Z_{t \wedge \tau_N}$?

Strong stationary times

Def A *strong stationary time* of a Markov process X is a randomized stopping time τ such that X_τ is independent of τ and X_τ is distributed according to its stationary law.

Claim Let Z be a Markov process with state space $\{0, 1, \dots, N\}$, started in $Z_0 = 0$, that jumps $k - 1 \mapsto k$ with rate $b_k > 0$ and $k \mapsto k - 1$ with rate $d_k > 0$ ($1 \leq k \leq N$). Then there exists a strong stationary time for Z and the fastest such time is distributed as a sum of independent exponentially distributed random variables whose parameters $\lambda_1 < \dots < \lambda_N$ are the negatives of the eigenvalues of the generator of Z .

Strong stationary times

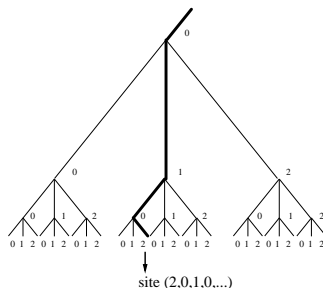


The hierarchical group

By definition, the *hierarchical group with freedom N* is the set

$$\Omega_N := \{i = (i_0, i_1, \dots) : i_k \in \{0, \dots, N-1\}, \\ i_k \neq 0 \text{ for finitely many } k\},$$

equipped with componentwise addition modulo N . Think of sites $i \in \Omega_N$ as the leaves of an infinite tree. Then i_0, i_1, i_2, \dots are the labels of the branches on the unique path from i to the root of the tree.

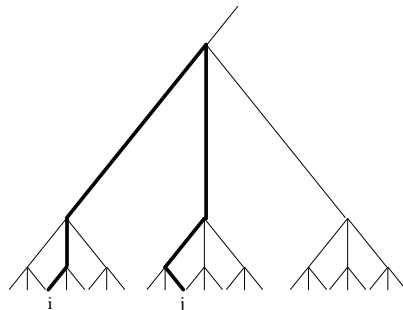


The hierarchical distance

Set

$$|i| := \inf\{k \geq 0 : i_m = 0 \ \forall m \geq k\} \quad (i \in \Omega_N).$$

Then $|i - j|$ is the *hierarchical distance* between two elements $i, j \in \Omega_N$. In the tree picture, $|i - j|$ measures how high we must go up the tree to find the last common ancestor of i and j .



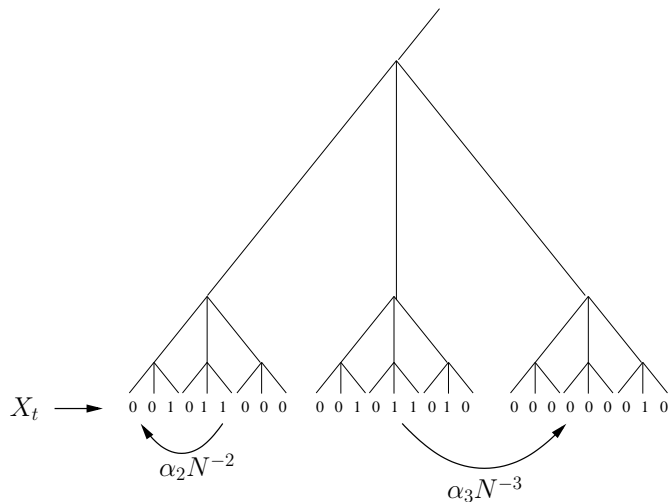
two sites at hierarchical distance 3

Hierarchical contact processes

Fix a *recovery rate* $\delta \geq 0$ and *infection rates* $\alpha_k \geq 0$ such that $\sum_{k=1}^{\infty} \alpha_k < \infty$. The *contact process* on Ω_N with these rates is the $\{0, 1\}^{\Omega_N}$ -valued Markov process $(X_t)_{t \geq 0}$ with the following description:

If $X_t(i) = 0$ (resp. $X_t(i) = 1$), then we say that the site $i \in \Omega_N$ is *healthy* (resp. *infected*) at time $t \geq 0$. An infected site i infects a healthy site j at hierarchical distance $k := |i - j|$ with rate $\alpha_k N^{-k}$, and infected sites become healthy with rate $\delta \geq 0$.

Hierarchical contact processes



Infection rates on the hierarchical group.

The critical recovery rate

We say that a contact process $(X_t)_{t \geq 0}$ on Ω_N with given recovery and infection rates *survives* if there is a positive probability that the process started with only one infected site never recovers completely, i.e., there are infected sites at any $t \geq 0$. For given infection rates, we let

$$\delta_c := \sup \left\{ \delta \geq 0 : \text{the contact process with infection rates } (\alpha_k)_{k \geq 1} \text{ and recovery rate } \delta \text{ survives} \right\}$$

denote the *critical recovery rate*. A simple monotone coupling argument shows that X survives for $\delta < \delta_c$ and dies out for $\delta > \delta_c$. It is not hard to show that $\delta_c < \infty$. The question whether $\delta_c > 0$ is more subtle.

(Non)triviality of the critical recovery rate

[Athreya & S. '10] Assume that $\alpha_k = e^{-\theta^k}$ ($k \geq 1$). Then:

(a) If $N < \theta$, then $\delta_c = 0$.

(b) If $1 < \theta < N$, then $\delta_c > 0$.

More generally, we show that $\delta_c = 0$ if

$$\liminf_{k \rightarrow \infty} N^{-k} \log(\beta_k) = -\infty, \quad \text{where} \quad \beta_k := \sum_{n=k}^{\infty} \alpha_n \quad (k \geq 1),$$

while $\delta_c > 0$ if

$$\sum_{k=m}^{\infty} (N')^{-k} \log(\alpha_k) > -\infty,$$

for some $m \geq 1$ and $N' < N$.

Proof of survival

We use added-on Markov processes to inductively derive bounds on the finite-time survival probability of finite systems. Let

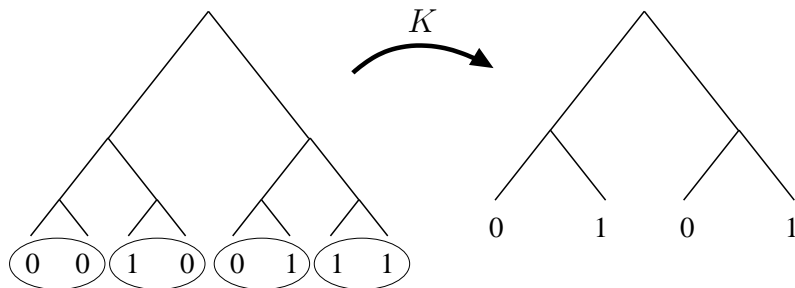
$$\Omega_2^n := \{i = (i_0, \dots, i_{n-1}) : i_k \in \{0, 1\}\}$$

and let $S_n := \{0, 1\}^{\Omega_2^n}$. We define a kernel from S_n to S_{n-1} by independently replacing blocks consisting of two spins by a single spin according to the stochastic rules:

$$\begin{aligned} 00 &\longrightarrow 0, & 11 &\longrightarrow 1, \\ \text{and } 01 \text{ or } 10 &\longrightarrow \begin{cases} 0 & \text{with probability } \xi, \\ 1 & \text{with probability } 1 - \xi, \end{cases} \end{aligned}$$

where $\xi \in (0, \frac{1}{2}]$ is a constant, to be determined later.

Renormalization kernel



The probability of this transition is $1 \cdot (1 - \xi) \cdot \xi \cdot 1$.

An added-on process

Let X be a contact process on Ω_2^n with infection rates $\alpha_1, \dots, \alpha_n$ and recover rate δ . Then X can be coupled to a process Y such that

$$\mathbb{P}[Y_t = y \mid (X_s)_{0 \leq s \leq t}] = K(X_t, y) \quad \text{a.s.} \quad (t \geq 0),$$

where K is the kernel defined before, and

$$\xi := \gamma - \sqrt{\gamma^2 - \frac{1}{2}} \quad \text{with} \quad \gamma := \frac{1}{4} \left(3 + \frac{\alpha_1}{2\delta} \right).$$

Moreover, the process Y can be coupled to a finite contact process Y' on Ω_2^{n-1} with recovery rate $\delta' := 2\xi\delta$ and infection rates $\alpha'_1, \dots, \alpha'_{n-1}$ given by $\alpha'_k := \frac{1}{2}\alpha_{k+1}$, in such a way that $Y'_t \leq Y_t$ for all $t \geq 0$.

Renormalization

We may view the map $(\delta, \alpha_1, \dots, \alpha_n) \mapsto (\delta', \alpha'_1, \dots, \alpha'_{n-1})$ as an (approximate) renormalization transformation. By iterating this map n times, we get a sequence of recovery rates $\delta, \delta', \delta'', \dots$, the last of which gives an upper bound on the spectral gap of the finite contact process X on Ω_2^n . Under suitable assumptions on the α_k 's, we can show that this spectral gap tends to zero as $n \rightarrow \infty$, and in fact, we can derive explicit lower bounds on the probability that finite systems survive till some fixed time t .

A question

Question Can we find an *exact* renormalization map $(\delta, \alpha_1, \dots, \alpha_n) \mapsto (\delta', \alpha'_1, \dots, \alpha'_{n-1})$ for hierarchical contact processes, i.e., for each hierarchical contact process X on Ω_2^n , can we find an *averaged Markov process* Y that is itself a hierarchical contact process on Ω_2^{n-1} ?